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Identifying Patient Subgroups with Treatment Effect Heterogeneity

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Background

Response Heterogeneity

Response Heterogeneity



Response Heterogeneity



Response Heterogeneity

- Data:
 - Population can be divided into subgroups
 - Investigated regimen might not be effective in all subgroups
 - Level of response might differ between subgroups
- Goal:
 - Identify for which subgroups the treatment is effective (response rate $\xi_k \ge x$ some threshold)

Aim

Factorial Design

Factorial Design

- Aim:
 - Identify patient subgroups with heterogeneity in APTs with factorial design
- Motivation:
 - For rheumatoid arthritis, combined therapies from various domains (*biologics/corticosteroids/nutritional supplements*): effective in some subgroups with different levels of effectiveness
 - Identify the predictors determining the effectiveness of combined therapies
- Example:
 - REMAP-CAP explored 240 (=5 antibiotics * 3 extended macrolides * 4 steroids * 4 antiviral therapies) potential regimens with a factorial design on different patient strata

Factorial Design



Approach

Bayesian Model Averaging

One regimen Full Model

- Suppose we only have 1 available treatment on 3 biomarker subgroups
- logistic regression with 1 categorical variable (biomarker group)
- k = 1,2,3

 $logit(\mathbb{E}(Y_{k,i})) = logit(\xi_k)$ $= \beta_0 + \beta_2 \mathbb{1}_2 + \beta_3 \mathbb{1}_3$



One regimen Full Model

- If we treat subgroups separately, the power might be low
- Allow for information sharing between subgroups
- Uncertainty about scope and strength of information sharing
- Psioda et al (2021) proposed Bayesian Model Averaging
 - Synthesizes information from different models
 - Compatible with adaptive design

One regimen Model Space



- K=3: Complete model space contains 5 models
- Each model partitions population into exchangeable sets
- For each set we estimate one common response rate

One regimen Model Space

 $logit(\mathbb{E}(Y_{k,i})) = logit(\xi_k)$ $= \beta_0 + \beta_2 \mathbb{1}_2 + \beta_3 \mathbb{1}_3$

$$\eta_2 - \eta_1 = \beta_2 = 0$$

$$\eta_3 - \eta_1 = \beta_3 = 0$$

$$\eta_3 - \eta_2 = \beta_3 - \beta_2 = 0$$

- 0 restrictions
 - Full model
 - No exchangeability
 - 1 possible way to select
- 1 restrictions
 - Reduced model
 - Pool 2 biomarker groups to be exchangeable
 - 3 possible way to select
- 2 restrictions
 - Minimal reduced model
 - Full exchangeability
 - 1 possible way to select

• 1+3+1=5 models

One regimen Model Averaging • Prior in the model space:

 $p(\mathcal{M}_m)$

- Posterior in the model space: $p(\mathcal{M}_m | \mathbf{D}) \propto p(\mathbf{D} | \mathcal{M}_m) p(\mathcal{M}_m)$
- Conditional on model m, posterior probability of efficiency: $P(\xi_k > x | \mathcal{M}_m, \mathbf{D})$
- Averaging across the model space, posterior probability of efficiency

$$P(\xi_k > x | \mathbf{D}) = \sum_m P(\xi_k > x | \mathcal{M}_m, \mathbf{D}) p(\mathcal{M}_m | \mathbf{D})$$

1*2 regimens

- Full Model: logistic regression
 - 2 categorical variables: regimen, biomarker
 - Main effect and interaction
- Regimen index j = 1,2
- Biomarker group index k = 1,2,3



$$\begin{split} logit(\mathbb{E}(Y_{(j)k,i})) &= logit(\xi_{(j)k}) \\ &= \beta_0 + \beta_2 \mathbb{1}_2 + \beta_3 \mathbb{1}_3 + \beta_{(2)} \mathbb{1}_{(2)} + \gamma_{(2)2} \mathbb{1}_{(2)2} + \gamma_{(2)3} \mathbb{1}_{(2)3} \end{split}$$

 $\mathbb{1}_{k'} = 1$ if the patient is in subgroup k'.

 $\mathbbm{1}_{(j')}=1$ if patient is randomized to accept regimen j'.

 $1\!\!1_{(j')k'} = 1\!\!1_{(j')} * 1\!\!1_{k'}.$

2*2 regimens

- logistic regression
 - 3 categorical variables
 - 2-way & 3-way interactions
- Regimens: (1,1) (1,2) (2,1) (2,2)
- Biomarker group index: k=1,2,3



$$\begin{split} logit(\mathbb{E}(Y_{(u,v)k,i})) &= logit(\xi_{(u,v)k}) \\ &= \beta_0 + \beta_2 \mathbb{1}_2 + \beta_3 \mathbb{1}_3 + \beta_{(\cdot,2)} \mathbb{1}_{(\cdot,2)} + \beta_{(2,\cdot)} \mathbb{1}_{(2,\cdot)} + \delta_{(2,2)} \mathbb{1}_{(2,2)} \\ &+ \kappa_{(2,\cdot)2} \mathbb{1}_{(2,\cdot)2} + \kappa_{(2,\cdot)3} \mathbb{1}_{(2,\cdot)3} + \kappa_{(\cdot,2)2} \mathbb{1}_{(\cdot,2)2} + \kappa_{(\cdot,2)3} \mathbb{1}_{(\cdot,2)3} \\ &+ \gamma_{(2,2)2} \mathbb{1}_{(2,2)2} + \gamma_{(2,2)3} \mathbb{1}_{(2,2)3} \end{split}$$

 $\mathbb{1}_{k'} = 1$ if the patient is in subgroup k'.

 $\mathbb{1}_{(u',\cdot)} = 1$ if patient is randomized to accept treatment u' from domain A. $\mathbb{1}_{(\cdot,v')} = 1$ if patient is randomized to accept treatment v' from domain B. $\mathbb{1}_{(u',\cdot)k'} = \mathbb{1}_{(u',\cdot)}\mathbb{1}_{k'}, \ \mathbb{1}_{(\cdot,v')k'} = \mathbb{1}_{(\cdot,v')}\mathbb{1}_{k'}, \ \mathbb{1}_{(u',v')k'} = \mathbb{1}_{k'}\mathbb{1}_{(u',\cdot)}\mathbb{1}_{(\cdot,v')}$

u = 1, 2; v = 1, 2; k = 1, 2, 3.

2*2 regimens Model Space

- (2*2 regimen) * 3 biomarker groups
- Each regimen cell has 5 possible partitioning configurations
- Complete model space contains 5⁴ models

$\eta_{(1,1)2} - \eta_{(1,1)1} = 0$	$\eta_{(1,2)2} - \eta_{(1,2)1} = 0$
$\eta_{(1,1)3} - \eta_{(1,1)1} = 0$	$\eta_{(1,2)3} - \eta_{(1,2)1} = 0$
$\eta_{(1,1)3} - \eta_{(1,1)2} = 0$	$\eta_{(1,2)3} - \eta_{(1,2)2} = 0$

$\eta_{(2,1)2} - \eta_{(2,1)1} = 0$	$\eta_{(2,2)2} - \eta_{(2,2)1} = 0$
$\eta_{(2,1)3} - \eta_{(2,1)1} = 0$	$\eta_{(2,2)3} - \eta_{(2,2)1} = 0$
$\eta_{(2,1)3} - \eta_{(2,1)2} = 0$	$\eta_{(2,2)3} - \eta_{(2,1)2} = 0$







Discussion

Challenges and Plans

Discussion

- BMA was proposed to borrow information dynamically, which performed well for binary (Psioda et al 2021) and continuous outcomes (Bean et al 2023)
 - Conjugate priors: Beta-binomial; Gamma-Normal
 - Averaged the complete model space
- Challenges to expand the method to factorial platform trials
 - No conjugate prior; factorial structure
 - Impossible to explore complete model space
- Plans
 - Used MCMC to estimate one factorial model
 - Next step: reduce computational burden; use specialized MCMC to sample from model space

Thank you

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References

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